# A tale of dual-approach construct validation and reliability testing for a *Zika* infection awareness knowledge questionnaire

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## Abstract

**Purpose** – This study aims to validate the English version of a WHO-adapted questionnaire: Zika infection awareness/knowledge questionnaire using a unique dual-approach validation model.

**Design/methodology/approach** – A cross-sectional pilot study of 30 adult respondents in Malaysia completed the self-administered questionnaire on knowledge and perception to Zika infection. Construct validity was assessed by exploratory factor analysis (EFA) of SPSS and Rasch partial credit. Reliability is tested using pKR20 and Cronbach's alpha.

**Findings** – Knowledge construct was unidimensional, good model fit, easy to endorse and well discriminative. Five-rating Likert scale for perception domain was appropriate. Knowledge domain should be separated into 6 level of difficulties. Perception domain should remain as one construct. Knowledge domain was highly reliability (pKR20 = 0.96), perception domain was fairly reliable (Cronbach's alpha = 0.641). Respondent's ability to answer knowledge domain and perception domain were separated into 3 and 4 levels.

**Research limitations/implications** – Small sample size may affect factor analysis.

**Practical implications** – The questionnaire has good psychometric properties to measure the knowledge and perception of Zika infection among Malaysian community.

**Social implications** – The questionnaire helped to gauge knowledge and perception of the general community in Malaysia to aid preparation of health education tool for Zika infection.

## Introduction

*Zika* infection has not attracted much attention in Malaysia despite being a country with an active and current circulation of the virus [1] and as the main contributor of Asian lineage of ZIKV [2]. Malaysia played an important role in the transmission of the *Zika* virus to Latin American regions in 2015 [3]. Underestimation is highly anticipated due to a suboptimal surveillance system [4].

Knowledge and perception can be used to predict the respondents' belief about their susceptibility toward the disease as well as the treatment of the disease according to the health belief model [5]. This maximizes the effort of prevention and control of Zika disease, especially in the design and strengthening of a capacity-building program such as the Community Behavior Intervention Program (COMBI) [6]. However, the majority of questionnaires were targeted at the clinical community and a bias for external validation of the community situation [7]. Sampling through social media or phone surveys may give rise to sampling bias, random bias, external validity and reliability issues [8]. The World Health Organization developed a questionnaire for communities with *Zika* virus transmission or those at risk such as Malaysia [9]. The resource pack provided general and thematic question banks in the domains of knowledge, attitudes and practices [9]. Users of this questionnaire can modify questions in the domains based on operational priorities [9]. The question style, language, answer mode and length can be modified for suitability of usage [9]. The questionnaire is available in various languages including English, Chinese, Portuguese, Arabic and Malay languages but no field validation was performed for knowledge/awareness toward *Zika* infection in South East Asia to the author's best knowledge [9].

This paper attempted to validate the construct validity and reliability of the English version of the WHO resource pack on knowledge, attitudes and practice surveys *Zika* virus disease and potential complications 2016 using a unique dual-approach validation model – the IBM SPSS and Rasch analysis. These two statistical approaches analyzed the data uniquely answering objectives, but the IBM SPSS was unable to measure underlying latent traits such as the respondents' ability, skills and language proficiency (item response theory) [10], which could be complemented by the Rasch, in the expression of item difficulty and individual ability and handling Likert scales category responses [11]. This formed the second unique feature of this paper – a bidirectional validation of a person's ability measured by scores achieved in the questionnaire and reversely validated the item ability of the questionnaire [12]. Third, this paper assessed both the construct validity and reliability of the questionnaire. With these three unique features, a solid validation was achieved to provide a valid and reliable questionnaire.

## Methodology

A cross-sectional pilot study of 30 adult respondents in Malaysia completed the questionnaire in a single attempt, assisted by trained enumerators from June to August 2018 via Google Forms. Categorization of sociodemographic characteristics of respondents was based on WHO classification of "elderly" for 60-year-olds and above [13]. Age was divided into "adult" and "elderly" groups as the author hypothesized that the knowledge and perception levels were different between both groups based on their place of residence and activities [14]. Education level was based on the local school system, which was later collapsed into high and low levels of education for distinctive comparison of knowledge.

*Instrument*
The questionnaire consisted of two domains – "knowledge" and "perception". The Knowledge domain had 14 questions adapted from section A of the General Question Bank of the WHO resource pack for *Zika* [9]. Items and answer option selections went through content validation with content experts in workshop discussion based on literature review and were made up of 63 items [9]. The letter "K" denoted "knowledge items", item KC-1 was on Information/communication, KS2-KS7 on cause/symptoms, KP8-KP9 on prevention, KT10-KT14 on treatment/care-seeking. The "perception" domain consisted of 6 items in a single construct denoted by the letter "P" and was developed by content experts. Perception of *vulnerability, transmission, complication, prevention, treatment, prevention* was in descending order of items (P1-6). This domain would be subjected to criterion validation through Likert category response in the Rasch [11].

Answer options for the "knowledge" domain were "Yes", "No" or "Don't know", in which only correct answers were given 1 mark, option "don't know" always carried no mark. Final scores were subjected to SPSS analysis to generate mean ± 2SD or median ± IQR. Scores were transformed into binary "1, 0" to generate logit data in model fitness statistics including item measure, person measure, mean ± 2SD score and reliability index.

The items in the "perception" domain were presented in a 5-point Likert scale where participants could answer "strongly agree", "agree", "neutral", "disagree" or "strongly disagree". Item P1, P2, P5 and P6 were tested on perception toward positive health-seeking behaviors, "strongly agree" was given a 5 score, "strongly disagree" was given 1 score. Items P3, P4 tested on perception toward a negative health-seeking behavior', scores were given in reverse order. Scores were inverted for Rasch analysis [11].

*Statistical analysis*
The author attempted a dual-approach of construct validation by marrying IBM SPSS and Rasch model analysis (Winsteps), which is rare but not the first of its kind [12]. Construct validity referred to the appropriateness of inferences made based on measurements to determine if a test measures the intended construct [15]. SPSS conventionally provides convergence and divergence of question items leading to construct formation, but was unable to measure the underlying ability and proficiency of respondents in the endorsement of responses to the question items tested [10]. The score respondents achieved may not accurately reflect their level of ability, as affected by how well-targeted the question items were including the language barrier and culture comprehensibility. The Rasch provided more comprehensive evidence on latent traits measured [12]. Besides, the Rasch was able to confirm the number of scales in measuring the difficulty in endorsing a response with the inference to a psychological attribute [16]. With this, the psychometric explanation was made clear and comparable with the range of item difficulty levels to the respondents' ability in answering the items [11].

*Difficulty index and discriminatory index*
Difficulty index was the proportion of participants who answered correctly, used to gauge the difficulty of questionnaire toward the respondents, calculated using the formula "$D$ = (students with correct answer × 100)/total students" in excel [15] and later compliment by the Rasch model (In summary statistic) [10].

Discriminatory index explained how well items discriminate different levels of traits observed, calculated by:

$$R = (H - L)/27\% \text{ of Total Respondents [10, 15]}$$

$H$ = number of correct answers from the top 27% of respondents

$L$ = number of correct answers from the bottom 27% of respondents

Discriminatory index of $\geq 0.4$ was considered very good items, 0.09–0.19 were indiscriminative items [10, 15].

Item-correlation measures relationship between items or discrete variables with equally-spaced values (e.g. Likert scale) [12, 15]. Good item-correlation ($r$ = 0.4–0.85) indicated model fitness, characterized by (1) Outfit MNSQ (outlier-sensitive fit) 0.5–1.5 logits, (2) Z-standard $-2$ to $+2.15$ Outfit MNSQ referred to mismatched ability to item difficulty or vice-versa [12]. Comparison of item correlation with "Corrected-Item-Correlation" using reliable analysis in SPSS complemented the test. Poor item-correlation ($r$ < 0.32 in Rasch and ITC < 0.3) were items poorly correlating difficulty levels in measuring respondents' ability to answer questions [10, 11].

*Reliability*
Reliability analysis is the reproducible ability of the questionnaire items to obtain the same result in different populations and/or times and was used as a form of criterion validation in this study. The internal consistency of measurements with dichotomous choices was calculated using pKR20, while those with Likert scales were calculated using Cronbach's alpha, where a value of >0.6 is desirable for both pKR20 and Cronbach's alpha [11]. The Rasch provides reliability analysis to both items and persons and is complemented with the SPSS. Confusing and/or problematic questions were then modified or excluded prior to the initiation of the actual study.

The number of scales or the quality of scale category rating was examined by a few criteria, that is (1) a total of at least $\geq 10$ respondents/category, (2) the rating frequencies of each category should be in regular distributions fashion (i.e. uniform, normal, bimodal or slightly skewed), (3) the average measures of rating should increase in uni-direction across the rating scale, (4) the total logit of adjacent threshold distance should be between 1.4 – 5. A logit >5 should suggest the merge of 2 adjacent scales, while logit <1.4 should suggest a split of scale into 2 or more adjacent categories, (d) a distinct probability curve displayed on each response category and (5) outfit value MNSQ <2 [11]. Distinctive peaks produced in probability curves (Rasch) indicated the number of rating suitable [11].

*Construct validation – exploratory factor analysis (SPSS) and dimensionality (Rasch)*
Principle component analysis aimed to reduce components to a smaller number by identifying principal components which explain the majority of variance [15]. Covariance with Varimax rotation was chosen to explore the 63 components (factors) of the knowledge domain and 6 factors in the perception domain. Criteria for principal components were 1) Eigenvalues above 1; 2) plots before the Scree plot levels off (Cattell rule); 3) when >80% variance explained [15]. Rasch item-dimensionality further analyzed the spread of variance and dimensionality of components by providing total variance % explained by measured trait and person' responses, besides items itself as a component [12]. Since the knowledge domain was built on 14 questions from four constructs in general question bank (A. Knowledge) WHO resource pack [9], expanded to 63 items by different answer options, the high possibility that a number of components are irreducible, as each option (e.g. sexual route transmission) was contributing significantly. In such a case, exploratory factor analysis (EFA) was used to confirm the validity of the 4 constructs [15].

Kaiser-Meyer-Olkin test of sampling adequacy and Bartlett's test of sphericity were used to determine the suitability of the data for factor analysis procedure [15]. However, the small sample size (<50) may be problematic for EFA, as generally, a 1:3 to 1:6 ratio per item should be available to extract principal components [17]. In the case of PCA not extractable, high

communalities (>0.7–0.9) of components (square root of variance) may explain the significance of the factor loading of the particular component, hence considered valid to assess the underlying trait [15]. A variance of >80% would also be adequate to indicate the adequacy of such a number of components in assessing the trait studied [15]. Factor loading less than 0.35 was suppressed for analysis. High factor loading, high item-correlation, high communalities indicate strong data despite near singular component due to small sample size [17].

*Model fitness statistic using Rasch analysis*
Model fitness was assessed by SPSS's factors' dimensionality, Item-correlation, discriminatory index and complemented by a summary statistic of Rasch [11]. Rasch partial credit rating affirmed the fitness of the Likert rating scale [11]. The fitness of items to trait measured would be expressed by good variance %, communality, strong factor loadings, good item-correlation and good discriminatory index in SPSS [10, 15]. While Rasch complemented the fitness of person toward items administered, represented by Infit MNSQ = 1.0 ($<-2$ to $<2.0$) and Infit ZSTD near 0 ($<-2$ to $<2.0$) [11]. Model fitness was also complemented by person reliability measure, items targetability (Wright map), adequacy of variance spread (dimensionality and Wright map), identify individual erratic performance (Infit MNSQ >2SD) and individual item-correlation [16]. All items within mean±1SD of person side (Wright map) were considered as questions items well-targeted to the participants, items >+1SD were too hard, <−1SD was too easy for a person. Erratic items should be revised or removed [11, 12].

*Ethical issue*
This research obtained ethical approval from the University Kebangsaan Malaysia with UKM/PPI/111/8/JEP-2019-104, National Medical Research Registration (NMRR) with the approval code NMRR-18-3853-45130-IIR and the Department of Welfare of Aborigine with code JAKOA/PP.30.032Jld 45(75).

**Results**
Sociodemographic characteristics of respondents are summarized as in Table 1. A total of 30 respondents from all over Malaysia completed the questionnaire, with 0% drop out, 70% were female respondents, 93.3% were Chinese, 66.7% were educated to university level, all respondents at least attained secondary education. 2/3 respondents were residing at the forest fringe, 40% were working in the forest fringe. The household size was small, with no pregnant respondents.

*Difficulty index versus person's ability – questions too hard or not understood?*
The mean score for knowledge was $61.9 \pm 24.76\%$; Rasch model analysis showed max mean measure for knowledge domain as log 18.6 (log 3.60 to -6.06), in descending order of ability. Questions that were most frequently missed were K4-6 "What causes *Zika*?" Breast milk (Answer: Yes); K5-3 "How does a person get *Zika*?" Through sexual intercourse (Answer: Yes); K13-3 "If a pregnant woman has *Zika*, what are the risks she faces?" She may have difficulty giving birth (Answer: Yes). All 3 questions were related to the sexual transmission of *Zika* infection, indicating a low level of knowledge toward this route of transmission for *Zika* virus, which is found uncommon among the *flavivirus* family.

The difficulty index of the knowledge domain of the questionnaire was 61.9 (61–80 = easy), complemented by Rasch to prove it as a set of well-targeted questions, 79.36% (50/63). The questions were not only easy for the Malaysian general public, but the items were easy to endorse with the local culture and language command. Beyond that, the Rasch Wright map showed

| Variables | | N | Percentage % |
|---|---|---|---|
| *Sociodemographic factors* | | | |
| *Age group (years)* | | | |
| | *Adult* | | |
| | 18–59 | 27 | 90 |
| | *Elderly* | | |
| | 60 and above | 3 | 10 |
| *Mean age (years) 33.17 ± 12.99 (kurtosis, −0.080, skewness 0.890)* | | | |
| *Sex* | | | |
| | Male | 9 | 30 |
| | Female | 21 | 70 |
| *Ethnicity* | | | |
| | Malay | 0 | 0 |
| | Chinese | 28 | 93.3 |
| | Indian | 2 | 6.7 |
| *Educational level* | | | |
| | *Low* | | *0* |
| | No formal school | | 0 |
| | Primary school | | 0 |
| | *High* | | *100* |
| | Secondary school | 10 | 33.3 |
| | Tertiary education | 20 | 66.7 |
| *Residence area* | | | |
| | Forest fringe | 21 | 70 |
| | City | 9 | 30 |
| *Work area* | | | |
| | Forest fringe | 12 | 40 |
| | City | 18 | 60 |
| Mean household number: 4.6 ± 1.673 (kurtosis 0.029, skewness 0.31) | | | |
| Mean household women at reproductive age (15-49 year old): 1.3 ± 1.393 (kurtosis 0.203, skewness 0.977) | | | |
| Median number of pregnant women in a household 0 ± 0.00 (kurtosis 23.77, skewness 4.782) | | | |
| Mean household monthly income (RM): 3000 + (25th centile 2000, 75th 6000) (kurtosis 2.62, skewness 8.958) | | | |
| Mean knowledge score: 39.0 ± 15.6/63 * 100% = 61.9 ± 24,76% (kurtosis, 1.318, skewness −1.418) | | | |

Table 1.
Sociodemographic characteristics of respondents

20.63% of the items <−1S.D. appeared too easy for the respondents. Participants could be separated into 3 groups: good/acceptable/poor knowledge based on the person separation index.

Difficulty index was not measured by SPSS for perception domain, but Rasch model affirmed 66.7% (4/6) of well-targeted question items, 33.3% of the items tested the ablest respondents. All respondents had good perceptions, in the singular person separation index.

*Discriminatory index – how well were question items targeted?*
The discriminatory index of the question items in the knowledge domain by SPSS was 0.34, indicating the items as "good items" to assess knowledge of the studied population on *Zika* infection. Rasch completed by showing items KS4-4, KS3-3, K3-5, KS2-1, KS5-7 and KT12 were items with poor correlation ($r < 0.32$) that caused indiscrimination of knowledge, similar to ITC (SPSS), KS4-2, KS4-6, KS5-4, KS6-6, KT10, KT12-3, KT13-3 ($r < 0.3$) (Table 2). A bubble chart identified two erratic items: KS4-2 and KT13-3. Item K4-2, "What causes *Zika*? Polluted water/ Dirty environment", the answer should be "No". Possible confusion may arise as pollution water means stagnant water bodies, which may encourage mosquito larvae breeding, leading to some attempts to answer "yes". K13-3, "If a pregnant woman has *Zika*, what are the risks she faces?" with the option "she may have difficulty giving birth". The answer to this question should be "Yes". The general population may not understand about anencephaly/microcephaly may be

| Knowledge domain | | | |
| SPSS | | Rasch | |
| --- | --- | --- | --- |
| Difficulty index (DI) | 61.9 (easy) | Wright map | 79.36% (50/63) items, well targeted<br>0% (0/63) items >+1SD, (too hard)<br>20.63% (13/63) <−1SD, (too easy) |
| Discriminatory index (Bi) | 0.34 (good items) | Fit statistic | Good model fitness item measure 0.96 (1.00)<br>Infit ZSTD −0.0 (closed to 0.00) |
| *Perception domain* | | | |
| SPSS | | Rasch | |
| Difficulty index (DI) | Not tested (Likert scale) | Wright map | 66.7% (4/6) items well targeted<br>33.3% (2/6) items >+1SD (too hard)<br>0% (0/6) <−1SD (too easy) |
| Discriminatory index (Bi) | 1.025 (very good items)<br>(Likert scale) | Fit statistic | Good model fitness item measure 0.98 (near 1.00)<br>Infit ZSTD = −0.2 (near 0.00) |

Table 2.
Difficulty index and discrimination index of knowledge and perception toward *Zika*

associated with polyhydramnios, unstable lie and hence unable to proceed with spontaneous vaginal delivery. Erratic questions should be revised, refined, or removed.

The perception domain had a DI of 1.025 and showed excellent discrimination. Person's ability ranged log 1.28 to–1.8. Two respondents (1.5%) show erratic performance (infit ZSTD 2.6). Only one perception item, P6 had a poor correlation with a person's ability in Rasch, $r = 0.24$ ($r < 0.32$), but three items were poorly correlated by SPSS (P1, P2, P6). Item P6, "I am confident that I can protect myself and my family from getting *Zika* infection", which "Strongly Agree" is given a 5 score. However, most respondents appeared not confident to protect themselves or family from getting *Zika* infection possibly due to fear of the disease, lack of knowledge of the disease and its prevention, over perception on the lethal nature of the disease likely due to sampling bias as researcher emphasis of the importance of the study to the studied population.

Rasch item separation index for knowledge domain was 6, items could be categorized into very easy, easy, moderate, above average, hard and very hard; and for perception, the domain was 4, which were easy, moderate, hard and very hard. Table 2 summarized the difficulty index and discriminatory index for both domains.

Perception domain response was assessed by the Likert scale. Rasch analysis enabled the Likert scale rating evaluation by checking on the probability curve for distinctive peaks (bond). Figure 1 shows the separation index yielded logit 2.72 (1.4 < s < 5.0), indicating a clear separation of 5 scales by distinctive intersections.

### Model fit by Rasch model analysis
Overall, both domains showed good model fitness to assess knowledge and risk perception of the studied population toward Zika disease. All question items for *Zika* infection knowledge and perception assessment had good person measure (Bi) and good item measure (Di). A bubble chart identified 2 erratic items for the knowledge domain, K4-2 & K13-3, perfect fitness for perception domain. This analysis completed discriminatory analysis bi-directionally, both on the items and the persons' abilities.

### Reliability
This questionnaire had very good reliability for the knowledge domain (pKR20 = 0.962) and fair reliability for the perception domain, Cronbach's alpha 0.641 with 4 items retained.

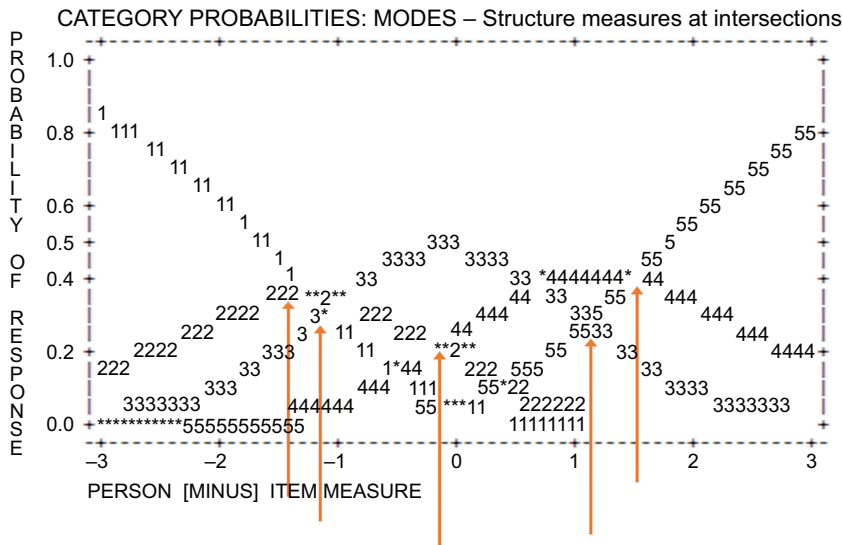CATEGORY PROBABILITIES: MODES – Structure measures at intersections



Figure 1.
Probability curves for
Likert scale in
perception domain

Cronbach's alpha was 0.418 for the initial 6 items in the perception domain. The value increased to 0.527 post removal of P2 and 0.641 with P2 & P6 removed. Question item P2 was "I will not get *Zika* infection if I practice safe sex, keep my environment clean and wear protective clothing" aimed to test on the perception of *Zika* disease transmission; while question item P6, "I am confident that I can protect myself and my family from getting *Zika* infection" aimed to test the perception of *Zika* disease prevention. In which an answer of "strongly agree" was given 5 points for both questions, and "strongly disagree" was given one point. The number of factor recovery was close to the expected constructs if all two problematic items were removed. Discussion with content experts on the revision of sentences and the number of items in the perception domain is needed to maintain the validity and reliability of the questionnaire. Table 3 summarizes the EFA and item analysis for both domains.

*Construct validation*
The Kaiser-Meyer-Olkin measure of sampling adequacy and Bartlett test of sphericity could not be generated and indicated inadequacy of sampling, most probably due to the small sample size or high multicollinearity issues. Component transformation matrix extracted four components which explained 54.66% of the variance, while 11 components were needed to achieve 81% variance (the above 80% rule) [15]. Scree plot levels off at 4 components, represented principal components, despite there being 15 components above Eigenvalue of 1 as described in Figure 2. A small sample size (<50) often gives rise to problematic near singular covariance matrix [17]. However, high commonality is seen in all items in knowledge (all above 0.7), indicating all components had significant loadings [15]. The standard error of measurement was 3.03 (<5), indicating that the finding is good (significant), Table 4 [15]. This offset the low variance with good psychometric properties [17]. This is most likely because the question items were all measuring the same trait – knowledge, while each item represented an option of the answer to the root question [18]. Therefore, all knowledge items should remain in one construct, four principal components explained nearly half of the variance.

Rasch item dimensionality affirmed that 38.7% of raw variance were explained by measure, 18.2% by items. (total up to 56.9%), while a person's ability to answer the question

| Items | Mean (S.D) | Factor 1 Information/ communication | Factor 2 Cause/ symptom | Factor 3 prevention | Factor 4 Treatment/ care-seeking | ITC |
|---|---|---|---|---|---|---|
| KC1 | 38.37 | 0.959 | | | | 0.609 |
| KS2 | 38.33 | | 0.960 | | | 0.518 |
| KS3-1 | 38.37 | | 0.959 | | | 0.623 |
| KS3-2 | 38.20 | | 0.959 | | | 0.578 |
| KS3-3 | 38.13 | | 0.959 | | | 0.736 |
| KS3-4 | 38.30 | | 0.959 | | | 0.708 |
| KS3-5 | 38.23 | | 0.959 | | | 0.698 |
| KS3-6 | 38.07 | | 0.959 | | | 0.730 |
| KS3-7 | 38.17 | | 0.959 | | | 0.738 |
| KS3-8 | 38.30 | | 0.959 | | | 0.665 |
| KS4-1 | 38.13 | | 0.960 | | | 0.555 |
| KS4-2 | 38.47 | | 0.961 | | | 0.009 |
| KS4-3 | 38.63 | | 0.960 | | | 0.372 |
| KS4-4 | 38.53 | | 0.960 | | | 0.546 |
| KS4-5 | 38.17 | | 0.959 | | | 0.575 |
| KS4-6 | 38.50 | | 0.961 | | | 0.269 |
| KS4-7 | 38.10 | | 0.960 | | | 0.443 |
| KS5-1 | 38.53 | | 0.959 | | | 0.705 |
| KS5-2 | 38.50 | | 0.960 | | | 0.487 |
| KS5-3 | 38.57 | | 0.960 | | | 0.434 |
| KS5-4 | 38.23 | | 0.961 | | | 0.247 |
| KS5-5 | 38.57 | | .960 | | | 0.428 |
| KS5-6 | 38.57 | | 0.960 | | | 0.302 |
| KS5-7 | 38.27 | | 0.960 | | | 0.434 |
| KS5-8 | 38.57 | | 0.960 | | | 0.509 |
| KS5-9 | 38.37 | | 0.960 | | | 0.516 |
| KS5-10 | 38.13 | | 0.960 | | | 0.494 |
| KS6-1 | 38.23 | | 0.959 | | | 0.736 |
| KS6-2 | 38.27 | | 0.959 | | | 0.601 |
| KS6-3 | 38.30 | | 0.959 | | | 0.666 |
| KS6-4 | 38.40 | | 0.959 | | | 0.564 |
| KS6-5 | 38.70 | | 0.960 | | | 0.469 |
| KS6-6 | 38.47 | | 0.961 | | | 0.240 |
| KS6-7 | 38.70 | | 0.960 | | | 0.545 |
| KS7 | 38.13 | | 0.960 | | | 0.310 |
| KP8 | 38.10 | | | 0.959 | | 0.742 |
| KP9-1 | 38.10 | | | 0.959 | | 0.705 |
| KP9-2 | 38.23 | | | 0.959 | | 0.705 |
| KP9-3 | 38.47 | | | 0.960 | | 0.489 |
| KP9-4 | 38.20 | | | 0.960 | | 0.411 |
| KP9-5 | 38.17 | | | 0.960 | | 0.546 |
| KP9-6 | 38.33 | | | 0.959 | | 0.569 |
| KP9-7 | 38.73 | | | 0.960 | | 0.387 |
| KT10 | 38.37 | | | | 0.961 | 0.066 |
| KT11-1 | 38.33 | | | | 0.959 | 0.725 |
| KT11-2 | 38.37 | | | | 0.959 | 0.621 |
| KT11-3 | 38.23 | | | | 0.959 | 0.665 |
| KT11-4 | 38.13 | | | | 0.959 | 0.713 |
| KT12-1 | 38.17 | | | | 0.959 | 0.711 |
| KT12-2 | 38.73 | | | | 0.959 | 0.614 |
| KT12-3 | 38.67 | | | | 0.961 | 0.222 |
| KT12-4 | 38.33 | | | | 0.960 | 0.389 |

Table 3.
Exploratory factor
analysis (knowledge
and perception of *Zika*)

(*continued*)

| Items | Mean (S.D) | Factor 1 Information/ communication | Factor 2 Cause/ symptom | Factor 3 prevention | Factor 4 Treatment/ care-seeking | | ITC | |
|---|---|---|---|---|---|---|---|---|
| KT12-5 | 38.47 | | | | 0.959 | | 0.569 | |
| KT12-6 | 38.30 | | | | 0.960 | | 0.545 | |
| KT13-1 | 38.23 | | | | 0.960 | | 0.435 | |
| KT13-2 | 38.50 | | | | 0.960 | | 0.463 | |
| KT13-3 | 38.23 | | | | 0.961 | | 0.110 | |
| KT14-1 | 38.27 | | | | 0.959 | | 0.601 | |
| KT14-2 | 38.33 | | | | 0.959 | | 0.745 | |
| KT14-3 | 38.33 | | | | 0.959 | | 0.564 | |
| KT14-4 | 38.27 | | | | 0.960 | | 0.550 | |
| KT14-5 | 38.27 | | | | 0.959 | | 0.686 | |
| KT14-6 | 38.10 | | | | 0.959 | | 0.627 | |
| *Cronbach's alpha* | *0.960* | *–* | *0.926* | *0.838* | *0.904* | | | |
| P1 | 13.17 | 6.63 | 0.341 | 0.252 | 0.254 | | 0.685 | |
| P2 | 12.47 | – | 0.527 | | −0.041 | | – | |
| P3 | 13.77 | 7.23 | 0.252 | 0.521 | 0.421 | | 0.514 | |
| P4 | 13.73 | 7.20 | 0.263 | 0.535 | 0.348 | | 0.482 | |
| P5 | 13.37 | 6.83 | 0.272 | 0.415 | 0.371 | | 0.578 | |
| P6 | 12.67 | – | 0.500 | | −0.053 | | – | |
| *Cronbach's alpha* | | | | | *0.418* | *0.641* | | |

Table 3.
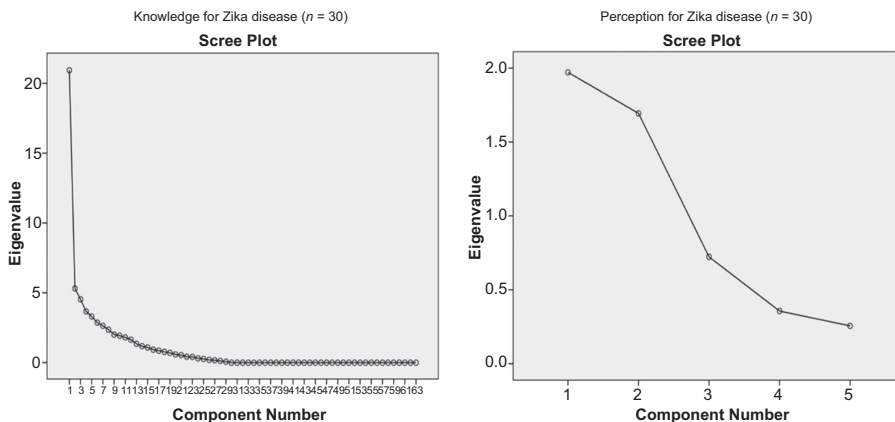


Figure 2.
Scree plots for knowledge & perception domains

(latent trait) explained 20.5% of the variance. Rasch analysis calculated construct separation by item separation index ($G$) using the formula:

$$G_{\text{item}} = \text{true SD}/\text{Real RMSE} = 1.46/0.37 = 3.95 = 4$$

Indicating four constructs were found to discriminate items at different levels of the trait being measured, in this case, "knowledge". The result compliments EFA in SPSS for construct validation.

For the perception domain, the Kaiser-Meyer-Olkin measure of sampling adequacy was 0.526, Bartlett's test of sphericity was significant at $p < 0.0001$, and the data were suitable

**Table 4.**
Construct validity for knowledge and perception domains using SPSS and Rasch analysis

| Summary statistic | Knowledge | Perception |
|---|---|---|
| Cronbach's $\alpha$ (Rasch) | 0.960 | 0.530 |
| Cronbach's $\alpha$ (SPSS) | 0.960 | 0.642 |
| Item reliability (Rasch) | 0.80 | 0.89 |
| Item separation (Rasch) | 4 | 3 |
| Person reliability (Rasch) | 0.960 | 0.50 |
| Person separation (Rasch) | 2 | 1 |
| Standard error (Rasch) | 0.36 | 0.27 |
| *Dimensionality* | | |
| Measured variance, % (Rasch) | 38.7% | 39.5% |
| Measured variance, % (SPSS) | 54.5% | 64.3% |
| Eigenvalue (Rasch) | 36.9 | 4 |
| Eigenvalue (SPSS) | 22.0 | 2 |
| Meaningful pattern of residuals (Rasch) | No | No |
| *Rating scale (Rasch)* | | |
| $\geq$10 responses per category | No | Yes |
| Regular distributions | N/A | Yes |
| Distance adjacent threshold | N/A | Yes |
| Outfit MNSQ <2 | Yes | Yes |
| Probability curves | 2 distinctive peaks | 5 distinctive peaks |
| *Number of construct (SPSS/Rasch)* | 4/4 | 2/3 |
| *Number of level trait measured (person's ability) (SPSS/Rasch)* | 3/3 | –/1 |
| *Number of constructs for item difficulty (Rasch)* | 6 | 4 |
| *Number of intersection* | – | 5 ratings Likert scale |

for factor analysis procedures despite the barely adequate sample size [17]. Communality was good, at 0.6–0.86, two components have Eigenvalues greater than 1, initial six-factor solution accounted for 64.3% variance. Total variance improved to 73.3% after problematic item P2 was removed. Pattern matrix showed strong factor loadings for all items into two principal components.

Rasch item dimensionality affirmed that 39.5% of raw variance were explained by measure, 29.6% by items; (total up to 56.9%), while a person's ability to answer the question (latent trait) explained 9.9% of the variance. Rasch analysis calculated construct separation by item separation index (G) using the formula:

$$G_{item} = \text{True SD}/\text{Real RMSE} = 0.577/0.22 = 2.59 = 3$$

Separating perception items to three constructs. The result compliments EFA in SPSS for construct validation (before rounded to nearest number, observe also scree plot pattern).

## Discussion

This study is a rare duet that assessed construct validity and reliability in a complementary and completing way bidirectionally [10, 12]. Validation with dual approaches provides extra value to validation of a questionnaire by the assessment of latent traits such as language barriers and subjects' cultural understanding toward a questionnaire design [12]. The questionnaire is suitable to be used in a large-scale study to assess the knowledge and perception of general communities in other developing countries in Southeast Asia with similar cultural understanding and language commands, to help governments tailor the capacity-building program in Zika disease prevention. The small sample size may serve as a limitation to factor analysis but does not leave researchers with no solution [17]. Beyond the rules-of-thumb of EFA ratio 5:1, a strong and uniform set of commonalities, the number of

items/factors and the item factor loading strengths all contribute to strong data [18] which produces reliable, reproducible, stable factor solution [19].

The questionnaire has good reliability and validity, good difficulty index & discriminatory index. The fair reliability of the perception domain is most probably related to inadequate health information received for *Zika* among the Malaysian community [15]. Erratic items KT13-3 & KS4-2 was possibly due to lack of clarity of phrases, wrong inference on the subject matter, or not testing the level of the respondents [15]. Respondents have good knowledge of *Zika*, similar to previous studies [7]. Nonetheless, the breakdown of construct analysis showed ineffective risk communication of *Zika* disease prevention among the population. There is a possibility of knowledge (on cause, symptom, prevention) inference toward *dengue* fever which presented similarly to *Zika* (KS2-KS7, KP8-KP9). Respondents failed to identify sexual transmission as a possible route (KS4-6, KS5-3, KS5-6, KS5-7, KS5-9). The vector prevention and control approach should be strengthened. Lack of knowledge on sexual and blood transmissions prevention potentially leads to devastating ZIKV endemicity [20]. However, positive health-seeking behaviors and good doctor-client relationship were found to be established (KT10-KT14), clients trusted the modality of treatment given by the ministry of health. This is important as the ZIKV disease monitoring period is much longer than that of *dengue* fever [21].

More than half of the respondents think that Malaysia is at risk of ZIKV, in contrast to research published [7] shortly after the Public Health Emergency of International Concern for *Zika* was declared in Malaysia. Most respondents had a good perception of *Zika* infection prevention and positive health-seeking behavior (P1-P4). Respondents' perceived risk of infection at forest fringe, thereby expressed keenness for self-protection. However, respondents expressed poor risk perception toward the threat & severity imposed by *Zika* infection (P5-P6), reflected respondents had not received adequate information on disease prevention from the local health authority, possibly overshadowed by the more common dengue fever and falsely assured by the undetected cases [4]. WHO risk communication guidelines should be implemented effectively to increase public resilience and participation in the collaborative effort to control the transmission of *Zika* [21].

## Conclusion
The field-validated English version *Zika* infection awareness/knowledge questionnaire had good psychometric properties to measure the knowledge and perception of *Zika* infection among the Malaysian community. Respondents had good knowledge and risk perception toward *Zika*. The questionnaire could aid the prevention and control of Zika disease in developing countries in Southeast Asia by designing a behavioral modification program and ensure effective risk communication. Further development should focus on the items revised for the Malaysian forest fringe population who are at higher risk of facing sylvatic to urban cycle transmission of *Zika* infection and should be tested on a larger scale.

Conflict of Interest: None

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